

Approximate Bayesian Computation: algorithms, theory and applications

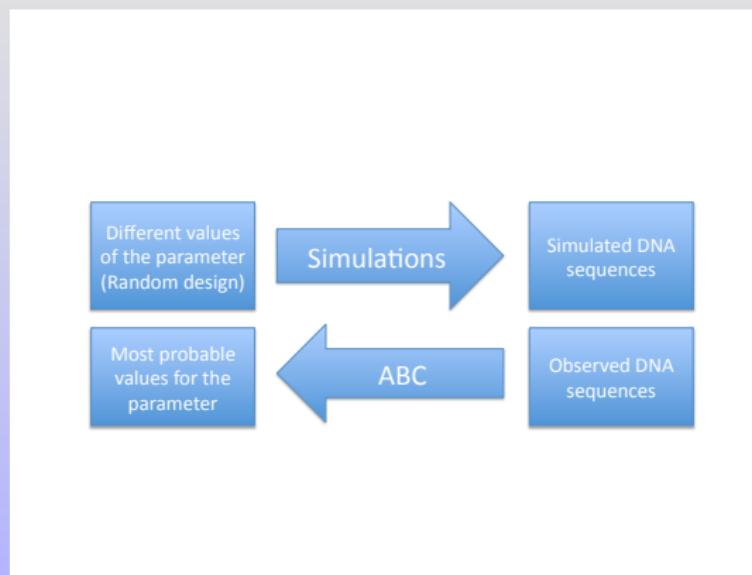
Michael G.B. Blum

Laboratoire TIMC-IMAG, UJF Grenoble, CNRS

MCEB, June 2012

"ABC is a 'democratizing' method in that it will attract, for example, biologists, who enjoy computer simulation but have little background in probability, into converting their favorite simulation into a tool for inference"

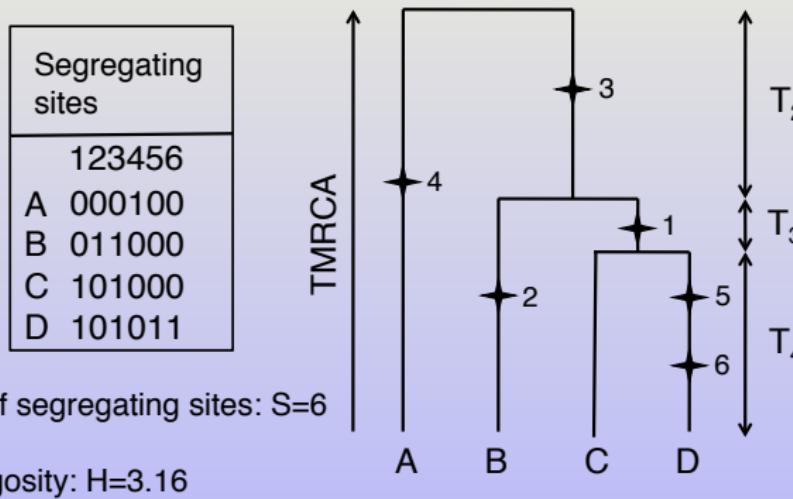
Beaumont and Rannala, Nat. Rev. Genet. 2004



Parameter inference for ABC

A coalescent example in population genetics

Estimating the mutation rate θ



Two approximations in ABC

- Replace *full posterior* $p(\theta|D)$ with *partial posterior* $p(\theta|\mathbf{s}_{obs})$
- Nonparametric estimation of $p(\theta|\mathbf{s}_{obs})$

Rejection algorithm

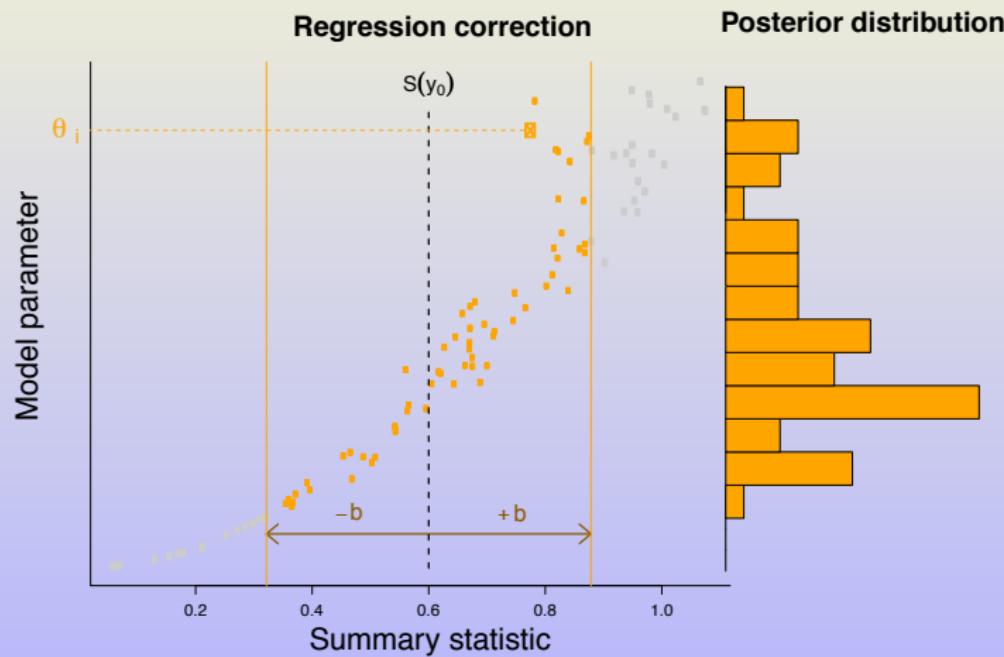
Pritchard et al., MBE 1999

- Simulate n values $\theta_i, i = 1, \dots, n$ from the prior π
- Simulate n (possibly multivariate) summary statistics \mathbf{s}_i according to $p(\mathbf{s}_i | \theta_i)$
- Consider the weighted sample $(\theta_i, W_i), i = 1, \dots, n$

$$W_i = \begin{cases} 1 & \text{if } \|\mathbf{s}_i - \mathbf{s}_{obs}\| \leq b \\ 0 & \text{otherwise.} \end{cases}$$

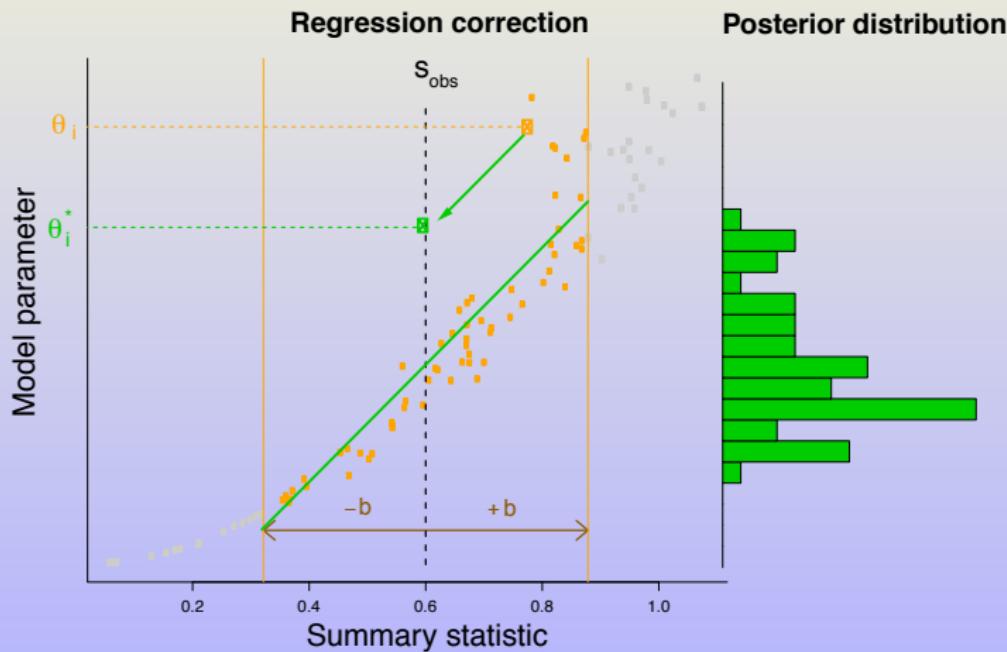
The parameter b is an acceptance threshold.

Rejection algorithm



Regression adjustment

Beaumont et al., Genetics 2002



Linear regression adjustment

- A model of local regression

$$\theta_i | \mathbf{s}_i = m(\mathbf{s}_i) + \epsilon_i$$

- Local linear approximation

$$m(\mathbf{s}_i) = \alpha + \mathbf{s}_i^t \boldsymbol{\beta}$$

- Adjustment

$$\theta_i^* = \hat{m}(\mathbf{s}_{obs}) + \tilde{\epsilon}_i,$$

Main theorem

Blum, JASA 2010

Asymptotic bias of the estimated posterior mean
 $j = 0$ rejection, $j = 1$ linear adjustment

$$C_{1,j} b^2$$

Asymptotic variance

$$\frac{C_3}{nb^d}$$

d is the number of the statistics and n is the number of simulations

Overemphasizes the curse of dimensionality because empirical evidence are much more optimistic.

Comparison between the two estimators with adjustment

When the model

$$\theta_i = m(\mathbf{s}_i) + \epsilon_i$$

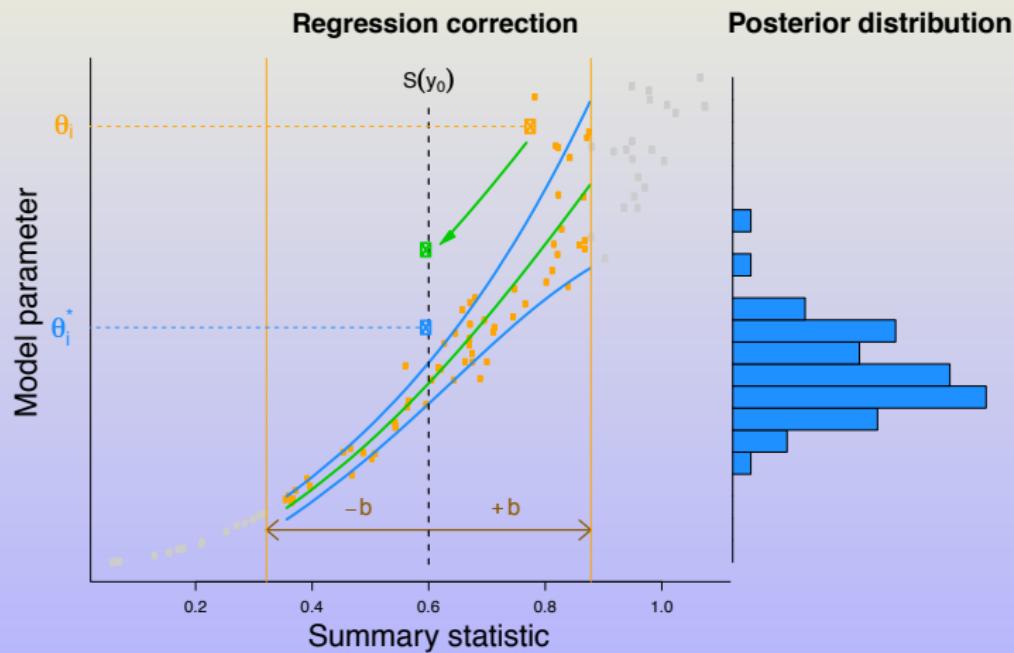
is homoscedastic in the vicinity of \mathbf{s}_{obs} , the bias for the estimator with quadratic adjustment is

$$o(b^2).$$

- Transformations of the sum stat to make the model as homoscedastic as possible.
- non-linear adj.
- Non-homoscedastic adjustment

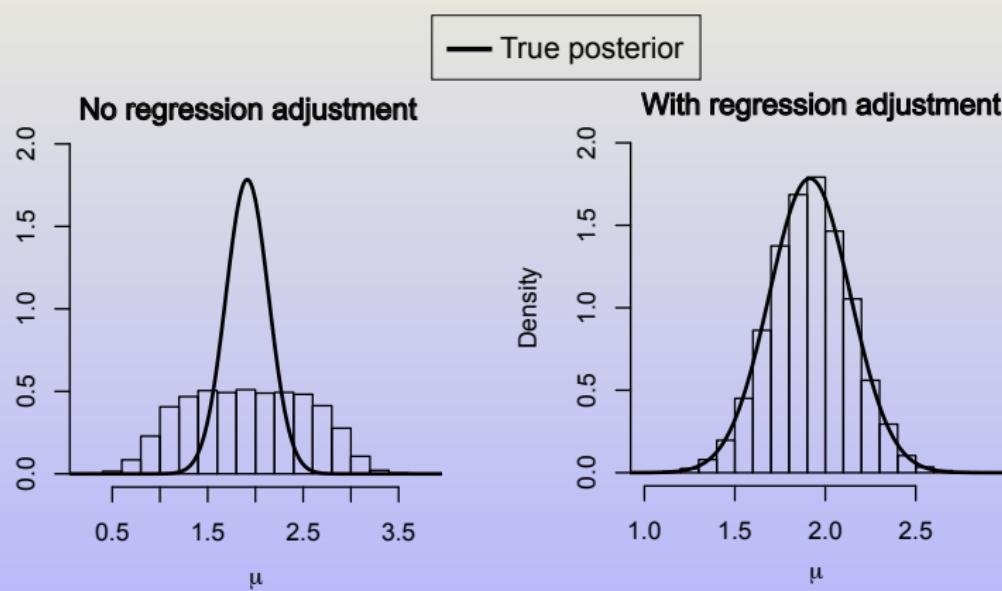
Regression adjustment for the mean and the variance

Blum and François, Stat and Comput 2010



ABC with and without adjustment

Estimating the mean in a Gaussian sample



How to check that ABC works when you do not know the posterior

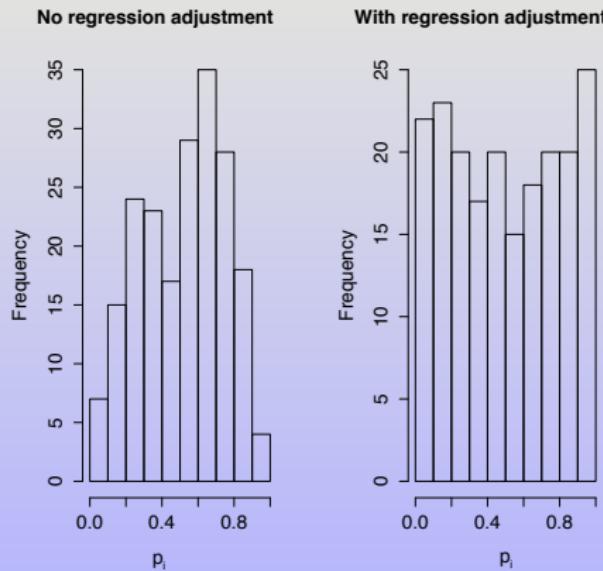
Cook et al. 2006 J. Comp. Graph. Stat.

- Take a (θ_i, \mathbf{s}_i) drawn from $\pi(\theta_i)p(\mathbf{s}_i|\theta_i)$.
- Perform ABC with $\mathbf{s}_{obs} = \mathbf{s}_i$.
- Compute the proportion p_i of posterior samples smaller than θ_i .

If the algorithm provide samples from $p(\theta_i|\mathbf{s}_i)$, p_i should be uniformly distributed.

How to check that ABC works when you do not know the posterior

Cook et al. 2006 J. Comp. Graph. Stat.



Adaptive ABC

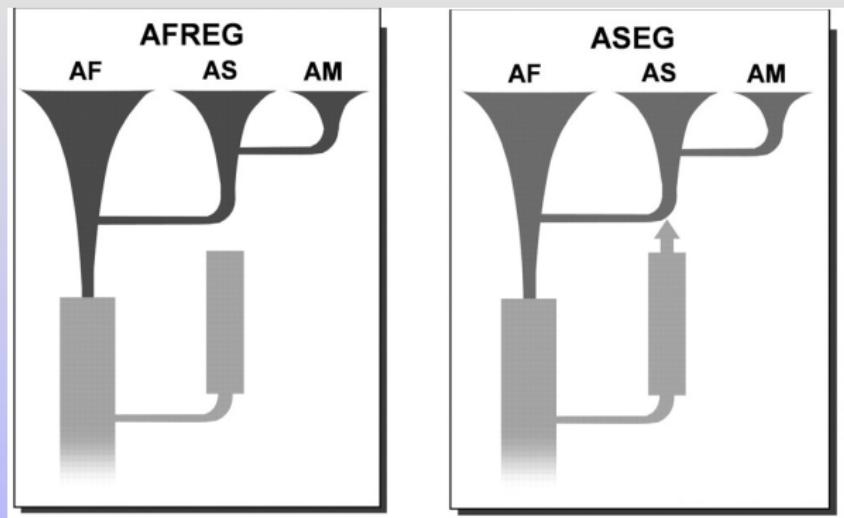
Sisson et al., PNAS 2007; Beaumont et al, Biometrika 2009; Del Moral et al. Stat and Comput 2011

Multi-step algorithms that sample θ from updated distributions that get closer and closer to the posterior distribution.

Model selection and related criticisms

Distinguishing between models An exemple in human evolution

Fagundes et al., PNAS 2007



Rejection algorithm

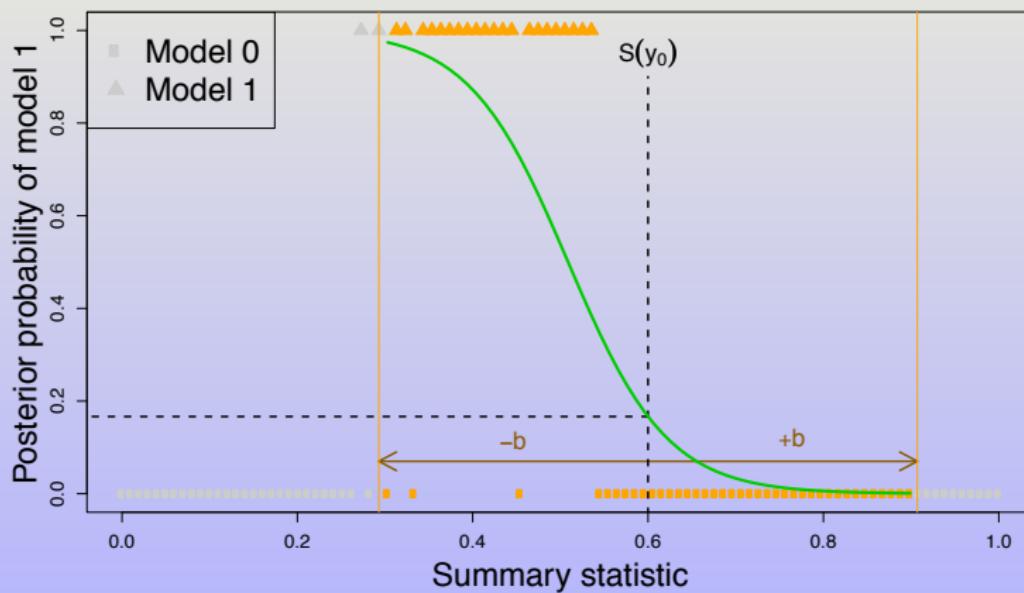
Pritchard et al., MBE 1999

- Simulate the same number of simulations under each model \mathcal{M}_k , $k = 1, \dots, K$.
- Accept the simulations for which $\|\mathbf{s}_i - \mathbf{s}_{obs}\| \leq b$.

The proportion of accepted simulations under each model $k = 1, \dots, K$ is an estimate of the posterior distribution $p(\mathcal{M}_k | \mathbf{s}_{obs})$.

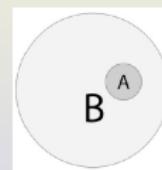
Model selection with logistic regression

Beaumont 2008



Criticisms about model selection with ABC

Templeton, PNAS 2010



'The probability of the nested special case must be less than or equal to the probability of the general model within which the special case is nested'.

- In a proper Bayesian framework, two hypotheses represent non-nested events.
- Coin tossing example. $\mathcal{M}_0 : q = 0.5$, $\mathcal{M}_1 : q \sim \mathcal{U}(0, 1)$.
9 heads out of 20 tosses.

$$p(\mathcal{M}_0 | 9/20) \approx 3.36 p(\mathcal{M}_1 | 9/20)$$

Criticisms about model selection with ABC

Robert et al., PNAS 2011

- ‘The algorithm involves an unknown loss of information induced by the use of insufficient summary statistics’.
- Assume that \mathbf{s} is sufficient for parameter inference in model 0 and model 1.

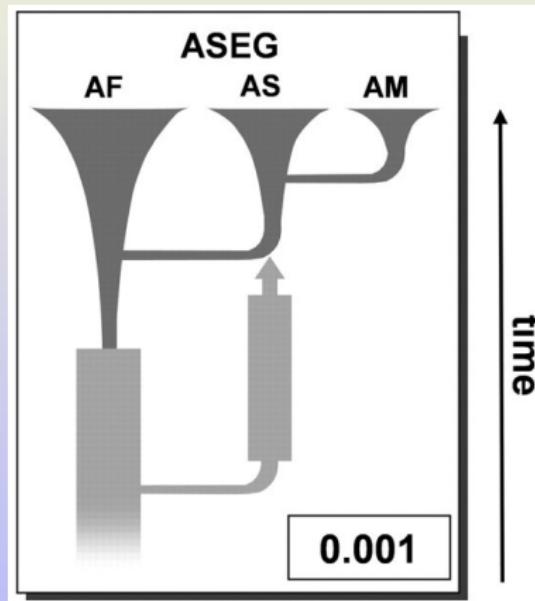
$$\frac{p(\mathcal{M}_0|\mathbf{s}_{obs})}{p(\mathcal{M}_1|\mathbf{s}_{obs})} = g(D) \frac{p(\mathcal{M}_0|D)}{p(\mathcal{M}_1|D)},$$

with $g(D)$ possibly different from 1.

Some elements of answers if your NSF reviewer keeps bothering you with this.

- The criticism pertains to situation where \mathbf{s} is sufficient for parameter inference in model 0 and model 1. Exercise : For coalescent models, try to find one for a constant-size population model and a bottleneck model.
- In approximate Bayesian Computation, we target $p(\mathcal{M}|\mathbf{s})$ instead of $p(\mathcal{M}|D)$.
- An important question to address might be ‘Does \mathbf{s} contain enough information to distinguish between models’.

Model selection with ABC : the right answer ?



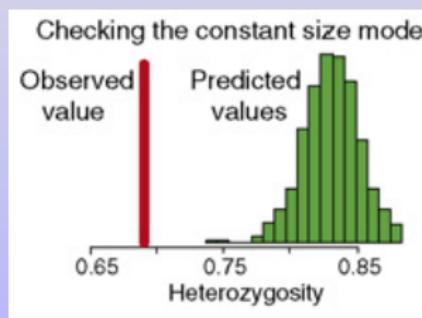
Fagundes et al., PNAS 2007

A deviance criterion for model selection with ABC

François and Laval, SAGMB 2011

Estimators of $p(\mathcal{M}|\mathbf{s})$ ignore regression adjustments on parameter samples

$$\text{DIC} = E_{\text{Post}}[\text{deviance}] + \text{effective number of parameters}$$



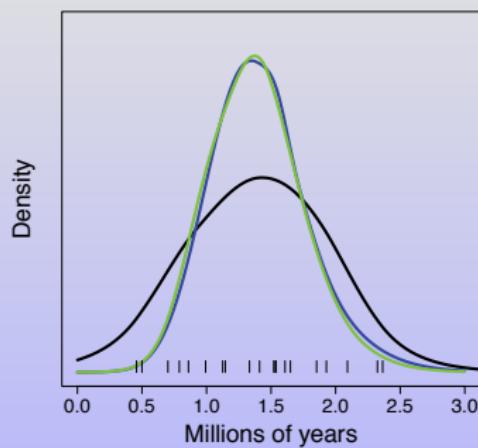
2-3 examples

Example 1 : Models of origins for modern humans

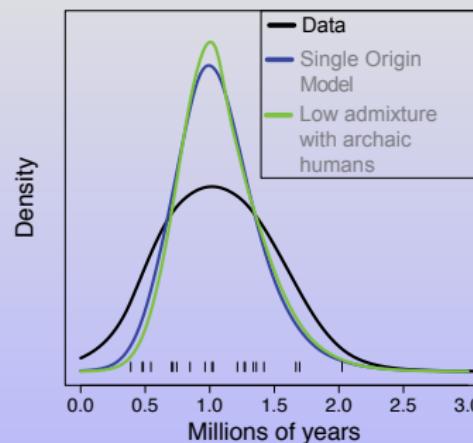
Blum and Jakobsson, MBE 2011

TMRCA distribution

A) Autosomal genes

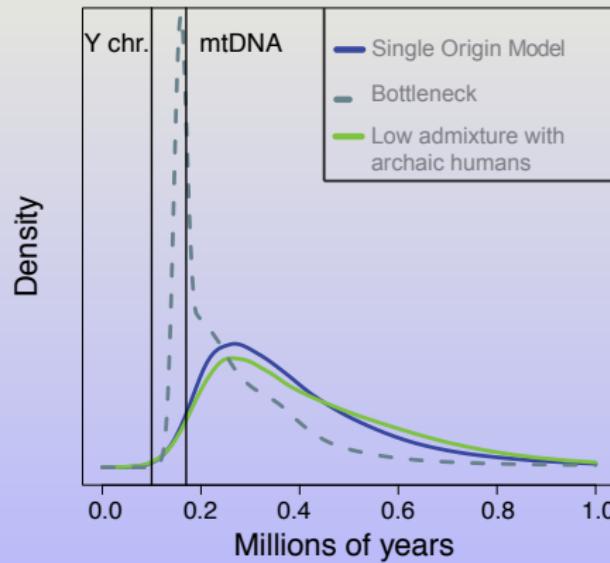


B) X chromosome



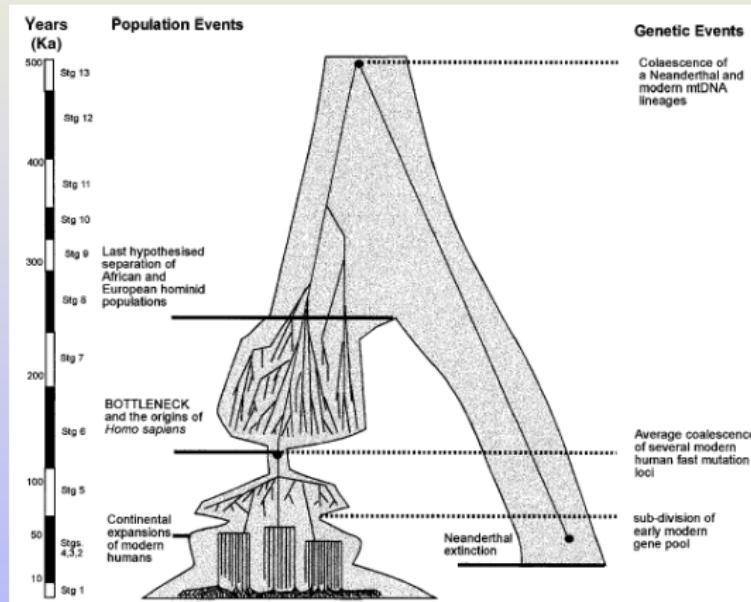
Models of origins for modern humans

C) mtDNA and Y chromosome



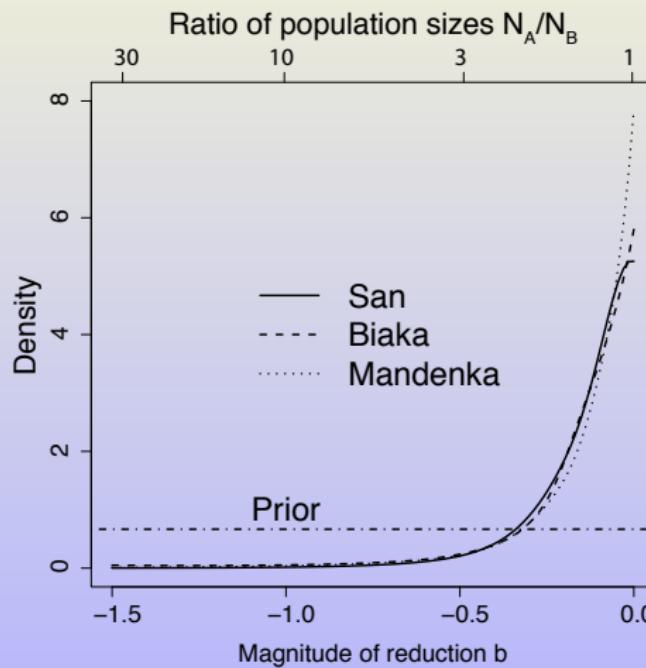
Example 1 : Testing the human 'speciation' bottleneck

Sjödin et al., MBE 2012



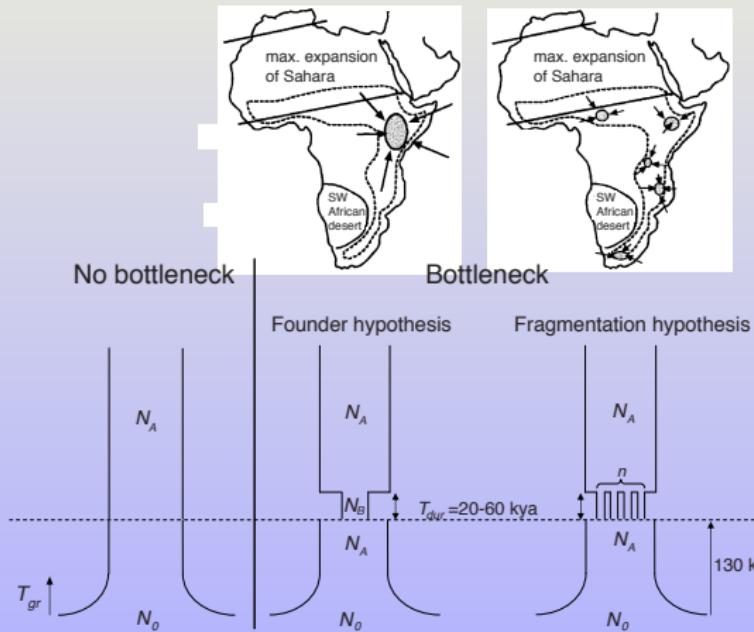
Lahr and Foley 1998

Estimating the strength of the bottleneck

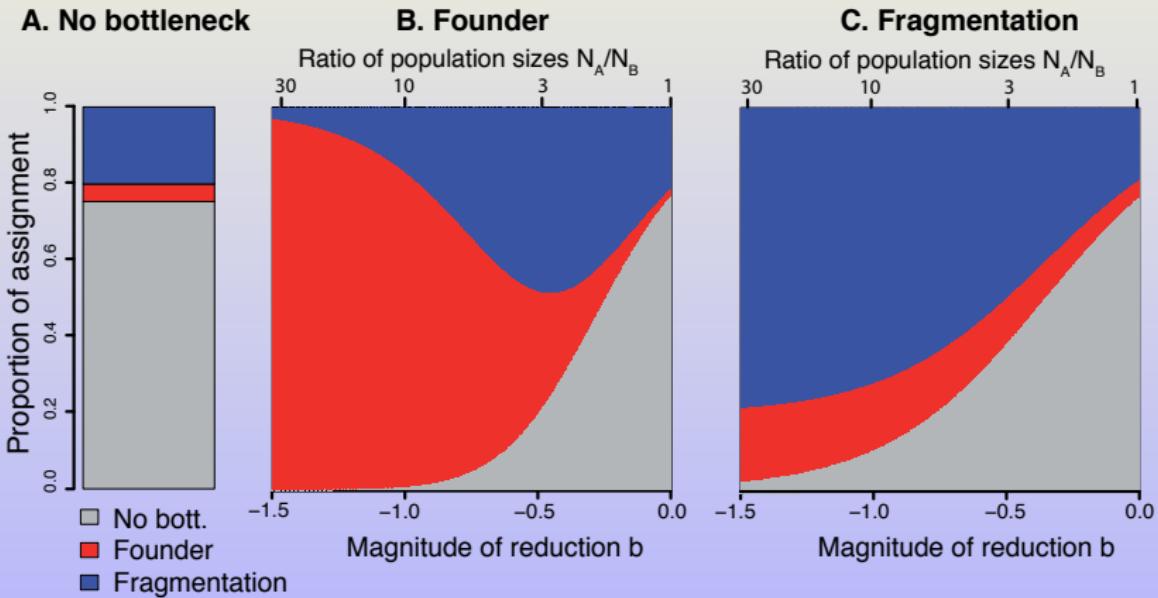


Support for a ‘no-bottleneck’ model against 2 bottleneck models

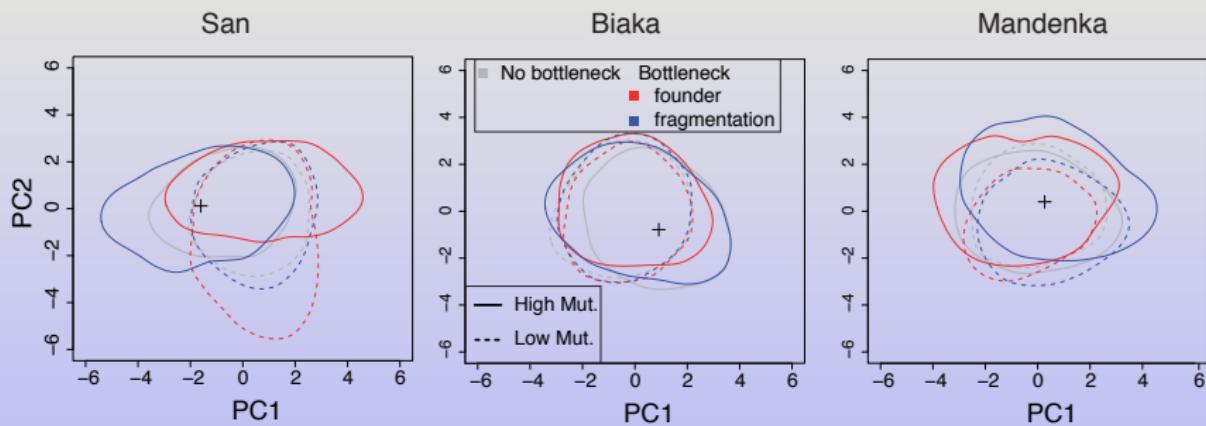
$Pr(\text{no bottleneck} | \mathbf{s}_{\text{obs}}) \geq 79\%$



Is it possible to distinguish between models ?

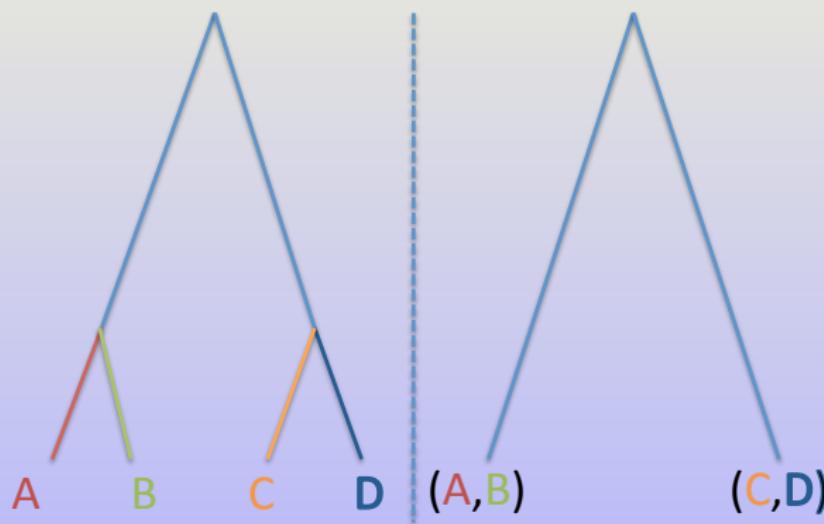


Goodness of fit Posterior predictive checks + PCA

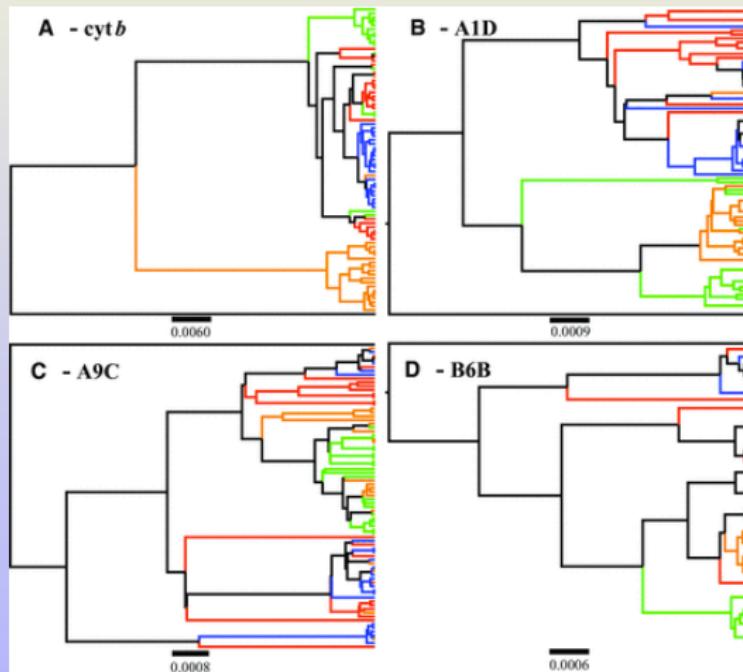


Example 2 : Species delimitation with ABC

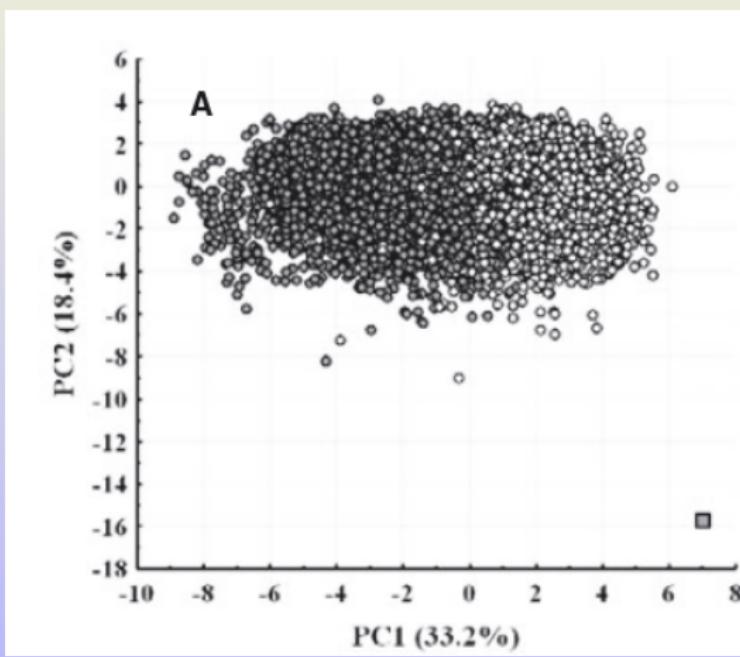
Camargo et al., Evolution 2012



Gene trees of loci sampled for species delimitation analyses



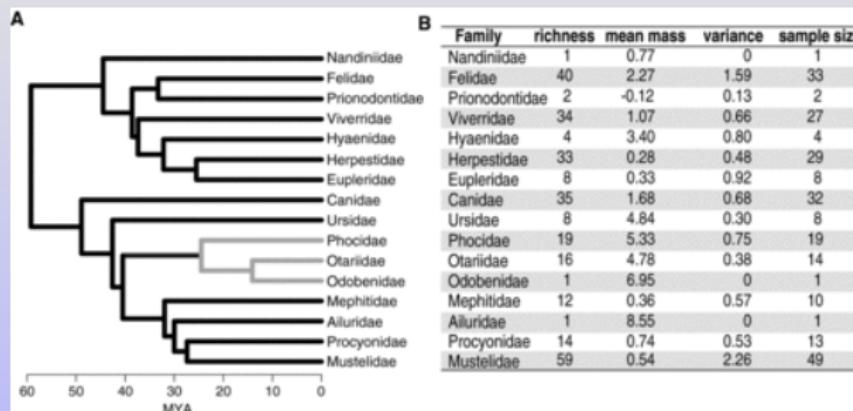
Prior predictive checks



Example 3 : Fitting models of continuous trait evolution

Slater et al., Evolution 2012

Time-calibrated phylogeny of Carnivora used to estimate rates of trait evolution

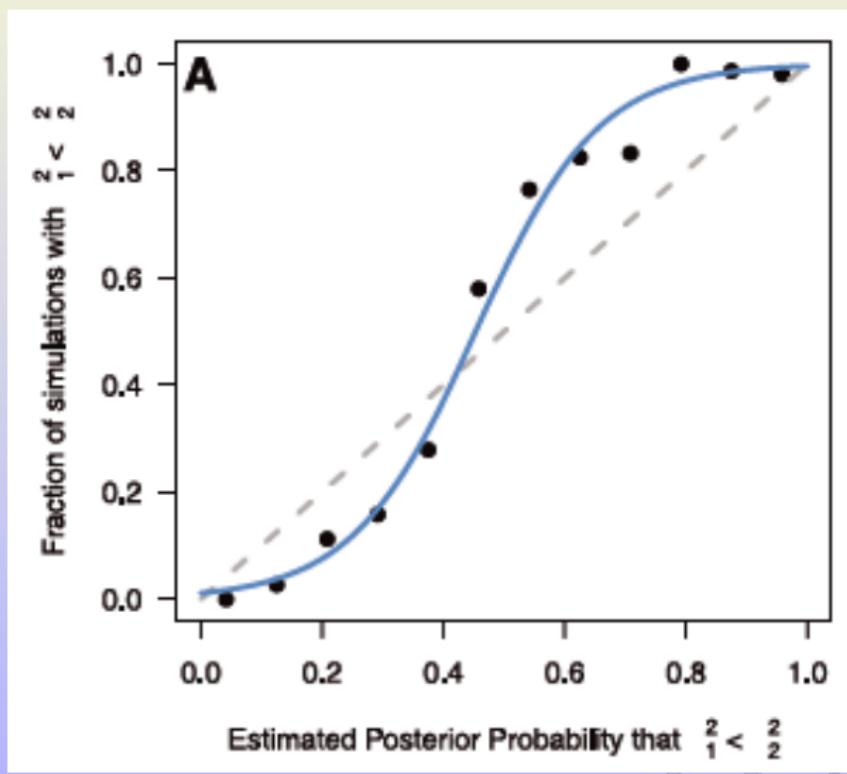


Comparing a two-rate model \mathcal{M}_2 to a one-rate model \mathcal{M}_1

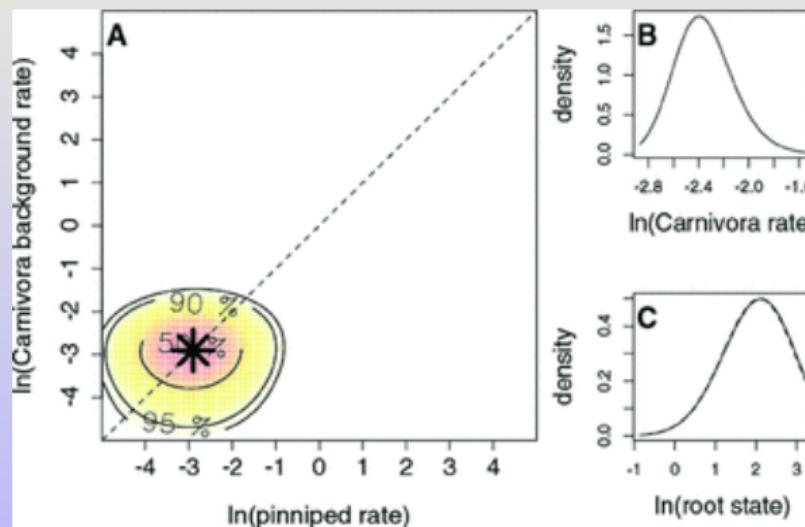
If $p(\mathcal{M}_1|\mathbf{s}) = x\%$, there is a probability of $x\%$ that \mathbf{s} was generated from \mathcal{M}_1 (and $1 - x\%$ that \mathbf{s} was generated from \mathcal{M}_2).

Cook et al. 2006 in a model selection framework.

Checking the ‘consistency’ of the Bayes factor



Comparing pinniped and terrestrial carnivore body size evolutionary rates



Conclusion

- ABC incorporates all aspects of Bayesian data analysis : formulation, fitting and model selection, and improvement of a model through model checking

Csilléry et al., TREE 2010.
- To address issues related to model selection
 - ➊ Ability to distinguish between models
 - ➋ ‘Consistency’ of the Bayes factor
 - ➌ Buy a bottle of wine to the reviewer
- The *R* package **abc** implements several ABC algorithms.

Colleagues



Olivier François



Mattias Jakobsson



Katalin Csilléry



Per Sjödin